# Logistic Regression

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Aim: Probability(Tomatoes | Radius) ? or



#### Idea: Use Linear Regression



 $P(X = Orange | Radius) = \theta_0 + \theta_1 \times Radius$ 

#### Idea: Use Linear Regression



$$P(X = Orange | Radius) = \theta_0 + \theta_1 \times Radius$$

Generally,

$$P(y=1|x)=X\theta$$

Prediction: If  $\theta_0 + \theta_1 \times Radius > 0.5 \rightarrow Orange$ Else  $\rightarrow$  Tomato Problem: Range of  $X\theta$  is  $(-\infty, \infty)$ But  $P(y = 1 | ...) \in [0, 1]$ 

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Linear regression for classification gives a poor prediction!

#### Ideal boundary



• Have a decision function similar to the above (but not so sharp and discontinuous)

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- Have a decision function similar to the above (but not so sharp and discontinuous)
- Aim: use linear regression still!

#### Idea: Use Linear Regression



Question. Can we still use Linear Regression? Answer. Yes! Transform  $\hat{y} \rightarrow [0, 1]$ 

### Logistic / Sigmoid Function

 $\hat{y} \in (-\infty, \infty)$   $\phi = \text{Sigmoid} / \text{Logistic Function} (\sigma)$  $\phi(\hat{y}) \in [0, 1]$ 



# Logistic / Sigmoid Function

 $Z 
ightarrow \infty$ 

 $z \to \infty$  $\sigma(z) \to 1$   $z \to \infty$   $\sigma(z) \to 1$  $z \to -\infty$   $z \to \infty$   $\sigma(z) \to 1$   $z \to -\infty$  $\sigma(z) \to 0$   $Z \to \infty$   $\sigma(Z) \to 1$   $Z \to -\infty$   $\sigma(Z) \to 0$ Z = 0  $z \to \infty$   $\sigma(z) \to 1$   $z \to -\infty$   $\sigma(z) \to 0$  z = 0 $\sigma(z) = 0.5$ 

#### Question. Could you use some other transformation ( $\phi$ ) of $\hat{y}$ s.t.

 $\phi(\hat{y}) \in [0,1]$ 

Yes! But Logistic Regression works.

$$P(y=1|X) = \sigma(X\theta) = \frac{1}{1+e^{-X\theta}}$$

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$$\therefore \frac{P(y=1|X)}{1-P(y=1|X)} = e^{X\theta} \implies X\theta = \log \frac{P(y=1|X)}{1-P(y=1|X)}$$

P(win) P(loss)

Here,

$$Odds = \frac{P(y = 1)}{P(y = 0)}$$
  
log-odds = log  $\frac{P(y=1)}{P(y=0)} = X\theta$ 

#### Q. What is decision boundary for Logistic Regression?

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or 
$$\frac{1}{1+e^{-X\theta}} = \frac{e^{-X\theta}}{1+e^{-X\theta}}$$
  
or  $e^{X\theta} = 1$   
or  $X\theta = 0$ 

Could we use cost function as:

$$J(\theta) = \sum (y_i - \hat{y}_i)^2$$
  
 $\hat{y}_i = \sigma(X\theta)$ 

Answer: No (Non-Convex) (See Jupyter Notebook)

#### Cost function convexity





Likelihood = 
$$P(D|\theta)$$
  
 $P(y|X,\theta) = \prod_{i=1}^{n} P(y_i|x_i,\theta)$   
where y = 0 or 1

Likelihood =  $P(D|\theta)$ 

$$P(y|X,\theta) = \prod_{i=1}^{n} P(y_i|x_i,\theta)$$
  
= 
$$\prod_{i=1}^{n} \left\{ \frac{1}{1+e^{-x_i^T \theta}} \right\}^{y_i} \left\{ 1 - \frac{1}{1+e^{-x_i^T \theta}} \right\}^{1-y_i}$$

[Above: Similar to  $P(D|\theta)$  for Linear Regression;

Difference Bernoulli instead of Gaussian]

 $-\log P(y|X, \theta) =$  Negative Log Likelihood = Cost function will be minimising  $= J(\theta)$ 

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- Idea find MLE estimate for heta

• 
$$p(H) = \theta$$
 and  $p(T) = 1 - \theta$ 

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- $\cdot \quad \frac{\partial \mathcal{LL}(\theta)}{\partial \theta} = 0 \implies \frac{n_h}{\theta} + \frac{n_t}{1-\theta} = 0 \implies \theta_{MLE} = \frac{n_h}{n_h + n_t}$

$$J(\theta) = -\log\left\{\prod_{i=1}^{n} \left\{\frac{1}{1+e^{-x_{i}^{T}\theta}}\right\}^{y_{i}} \left\{1-\frac{1}{1+e^{-x_{i}^{T}\theta}}\right\}^{1-y_{i}}\right\}$$
$$J(\theta) = -\left\{\sum_{i=1}^{n} y_{i} \log(\sigma_{\theta}(x_{i})) + (1-y_{i}) \log(1-\sigma_{\theta}(x_{i}))\right\}$$

$$\begin{aligned} \frac{\partial J(\theta)}{\partial \theta_j} &= -\frac{\partial}{\partial \theta_j} \bigg\{ \sum_{i=1}^n y_i \log(\sigma_\theta(x_i)) + (1 - y_i) \log(1 - \sigma_\theta(x_i)) \bigg\} \\ &= -\sum_{i=1}^n \bigg[ y_i \frac{\partial}{\partial \theta_j} \log(\sigma_\theta(x_i)) + (1 - y_i) \frac{\partial}{\partial \theta_j} \log(1 - \sigma_\theta(x_i)) \bigg] \end{aligned}$$

$$\frac{\partial J(\theta)}{\partial \theta_j} = -\sum_{i=1}^n \left[ y_i \frac{\partial}{\partial \theta_j} \log(\sigma_{\theta}(x_i)) + (1 - y_i) \frac{\partial}{\partial \theta_j} \log(1 - \sigma_{\theta}(x_i)) \right]$$

$$= -\sum_{i=1}^{n} \left[ \frac{y_i}{\sigma_{\theta}(x_i)} \frac{\partial}{\partial \theta_j} \sigma_{\theta}(x_i) + \frac{1 - y_i}{1 - \sigma_{\theta}(x_i)} \frac{\partial}{\partial \theta_j} (1 - \sigma_{\theta}(x_i)) \right]$$
(1)

Aside:

$$\frac{\partial}{\partial z}\sigma(z) = \frac{\partial}{\partial z}\frac{1}{1+e^{-z}} = -(1+e^{-z})^{-2}\frac{\partial}{\partial z}(1+e^{-z})$$
$$= \frac{e^{-z}}{(1+e^{-z})^2} = \left(\frac{1}{1+e^{-z}}\right)\left(\frac{e^{-z}}{1+e^{-z}}\right) = \sigma(z)\left\{\frac{1+e^{-z}}{1+e^{-z}} - \frac{1}{1+e^{-z}}\right\}$$
$$= \sigma(z)(1-\sigma(z))$$

#### Resuming from (1)

$$\begin{split} \frac{\partial J(\theta)}{\partial \theta_j} &= -\sum_{i=1}^n \left[ \frac{y_i}{\sigma_\theta(x_i)} \frac{\partial}{\partial \theta_j} \sigma_\theta(x_i) + \frac{1 - y_i}{1 - \sigma_\theta(x_i)} \frac{\partial}{\partial \theta_j} (1 - \sigma_\theta(x_i)) \right] \\ &= -\sum_{i=1}^n \left[ \frac{y_i \sigma_\theta(x_i)}{\sigma_\theta(x_i)} (1 - \sigma_\theta(x_i)) \frac{\partial}{\partial \theta_j} (x_i \theta) + \frac{1 - y_i}{1 - \sigma_\theta(x_i)} (1 - \sigma_\theta(x_i)) \frac{\partial}{\partial \theta_j} (1 - \sigma_\theta(x_i)) \right] \\ &= -\sum_{i=1}^n \left[ y_i (1 - \sigma_\theta(x_i)) x_i^j - (1 - y_i) \sigma_\theta(x_i) x_i^j \right] \\ &= -\sum_{i=1}^n \left[ (y_i - y_i \sigma_\theta(x_i) - \sigma_\theta(x_i) + y_i \sigma_\theta(x_i)) x_i^j \right] \\ &= \sum_{i=1}^n \left[ \sigma_\theta(x_i) - y_i \right] x_i^j \end{split}$$

$$rac{\partial J(\theta)}{\theta_j} = \sum_{i=1}^N \left[ \sigma_{\theta}(x_i) - y_i \right] x_i^j$$

Now, just use Gradient Descent!

#### Cost function convexity





The Hessian matrix of f(.) with respect to  $\theta$ , written  $\nabla_{\theta}^2 f(\theta)$  or simply as  $\mathbb{H}$ , is the  $d \times d$  matrix of partial derivatives,

$ abla^2_ heta f( heta) =$	$\begin{bmatrix} \frac{\partial^2 f(\theta)}{\partial \theta_1^2} \\ \frac{\partial^2 f(\theta)}{\partial \theta_1^2} \end{bmatrix}$	$rac{\partial^2 f(\theta)}{\partial \theta_1 \partial \theta_2}$ $\partial^2 f(\theta)$		$\frac{\partial^2 f(\theta)}{\partial \theta_1 \partial \theta_n} = \frac{\partial^2 f(\theta)}{\partial \theta_1 \partial \theta_n}$
	$\frac{\partial}{\partial \theta_2 \partial \theta_1}$	$\frac{\partial \theta_1(\theta)}{\partial \theta_2^2}$	•••	$\frac{\partial \theta_2}{\partial \theta_2} \frac{\partial \theta_n}{\partial \theta_n}$
			•••	
		• • •	• • •	• • •
	$\frac{\partial^2 f(\theta)}{\partial \theta_n \partial \theta_1}$	$rac{\partial^2 f( heta)}{\partial  heta_n \partial  heta_2}$		$\frac{\partial^2 f(\theta)}{\partial \theta_p^2}$

The most basic second-order optimization algorithm is Newton's algorithm, which consists of updates of the form,

$$\theta_{k+1} = \theta_k - \mathbb{H}_k^1 g_k$$

where  $g_k$  is the gradient at step k. This algorithm is derived by making a second-order Taylor series approximation of  $f(\theta)$  around  $\theta_k$ :

$$f_{quad}(\theta) = f(\theta_k) + g_k^T(\theta - \theta_k) + \frac{1}{2}(\theta - \theta_k)^T \mathbb{H}_k(\theta - \theta_k)$$

differentiating and equating to zero to solve for  $\theta_{k+1}$ .

Now assume:

$$g(\theta) = \sum_{i=1}^{n} \left[ \sigma_{\theta}(\mathbf{x}_{i}) - \mathbf{y}_{i} \right] \mathbf{x}_{i}^{j} = \mathbf{X}^{\mathsf{T}} (\sigma_{\theta}(\mathbf{X}) - \mathbf{y})$$
$$\pi_{i} = \sigma_{\theta}(\mathbf{x}_{i})$$

Let  $\mathbb{H}$  represent the Hessian of  $J(\theta)$ 

$$\mathbb{H} = \frac{\partial}{\partial \theta} g(\theta) = \frac{\partial}{\partial \theta} \sum_{i=1}^{n} \left[ \sigma_{\theta}(x_{i}) - y_{i} \right] x_{i}^{j}$$
$$= \sum_{i=1}^{n} \left[ \frac{\partial}{\partial \theta} \sigma_{\theta}(x_{i}) x_{i}^{j} - \frac{\partial}{\partial \theta} y_{i} x_{i}^{j} \right]$$
$$= \sum_{i=1}^{n} \sigma_{\theta}(x_{i}) (1 - \sigma_{\theta}(x_{i})) x_{i} x_{i}^{T}$$
$$= \mathbf{X}^{\mathsf{T}} diag(\sigma_{\theta}(x_{i}) (1 - \sigma_{\theta}(x_{i}))) \mathbf{X}_{i}$$

### Iteratively reweighted least squares (IRLS)

For binary logistic regression, recall that the gradient and Hessian of the negative log-likelihood are given by:

$$g(\theta)_{k} = \mathbf{X}^{\mathsf{T}}(\pi_{\mathsf{k}} - \mathsf{y})$$
  

$$\mathbf{H}_{k} = \mathbf{X}^{\mathsf{T}}S_{k}\mathbf{X}$$
  

$$\mathbf{S}_{k} = diag(\pi_{1k}(1 - \pi_{1k}), \dots, \pi_{nk}(1 - \pi_{nk}))$$
  

$$\pi_{ik} = sigm(\mathbf{x}_{i}\theta_{\mathsf{k}})$$

The Newton update at iteraion k + 1 for this model is as follows:

$$\begin{aligned} \theta_{k+1} &= \theta_k - \mathbb{H}^{-1} g_k \\ &= \theta_k + (X^T S_k X)^{-1} X^T (y - \pi_k) \\ &= (X^T S_k X)^{-1} [(X^T S_k X) \theta_k + X^T (y - \pi_k)] \\ &= (X^T S_k X)^{-1} X^T [S_k X \theta_k + y - \pi_k] \end{aligned}$$

Unregularised:

$$J_1(\theta) = -\left\{\sum_{i=1}^n y_i \log(\sigma_{\theta}(x_i)) + (1 - y_i) \log(1 - \sigma_{\theta}(x_i))\right\}$$

L2 Regularization:

$$J(\theta) = J_1(\theta) + \lambda \theta^{\mathsf{T}} \theta$$

L1 Regularization:

$$J(\theta) = J_1(\theta) + \lambda |\theta|$$

- 1. Use one-vs.-all on Binary Logistic Regression
- 2. Use one-vs.-one on Binary Logistic Regression
- 3. Extend <u>Binary</u> Logistic Regression to <u>Multi-Class</u> Logistic Regression

$$Z \in \mathbb{R}^d$$
$$f(z_i) = \frac{e^{z_i}}{\sum_{i=1}^d e^{z_i}}$$
$$\therefore \sum f(z_i) = 1$$

 $f(z_i)$  refers to probability of class <u>i</u>

$$k = 1, \dots, k$$
classes  
 $P(y = k | x, \theta) = \frac{e^{x\theta_k}}{\sum_{k=1}^{K} e^{x\theta_k}}$ 

#### Softmax for Multi-Class Logistic Regression

For K = 2 classes,

$$P(y = k | x, \theta) = \frac{e^{x\theta_k}}{\sum_{k=1}^{K} e^{x\theta_k}}$$
$$P(y = 0 | x, \theta) = \frac{e^{x\theta_0}}{e^{x\theta_0} + e^{x\theta_1}}$$
$$P(y = 1 | x, \theta) = \frac{e^{x\theta_1}}{e^{x\theta_0} + e^{x\theta_1}} = \frac{e^{x\theta_1}}{e^{x\theta_1}\{1 + e^{x(\theta_0 - \theta_1)}\}}$$
$$= \frac{1}{1 + e^{-x\theta'}}$$
$$= \text{Sigmoid!}$$

For 2 class we had:

$$J(\theta) = -\left\{\sum_{i=1}^{n} y_i \log(\sigma_{\theta}(x_i)) + (1 - y_i) \log(1 - \sigma_{\theta}(x_i))\right\}$$

Extend to K-class:

$$J(\theta) = \left\{ \sum_{i=1}^{n} \sum_{k=1}^{K} I\{y_i = k\} \log \frac{e^{x_i \theta_k}}{\sum_{k=1}^{K} e^{x_i \theta_k}} \right\}$$
$$i \to \text{Sample #} \qquad \text{I: Identity Function}$$
$$k \to Class \qquad I(\text{true}) = 1; I(\text{false}) = 0$$

Now:

$$\frac{\partial J(\theta)}{\partial \theta_k} = \sum_{i=1}^n \left[ x_i \left\{ l(y_i = k) - P(y_i = k | x_i, \theta) \right\} \right]$$